

# Small Area Estimation under the Growth Curve model

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# Outline

- Introduction
- The model formulation
- Estimation of model parameters
- Prediction of random effects
- Simulation study example
- Further research
- Some references

# Introduction

- The term *Growth Curve Modeling* has been used in different contexts to refer to a wide array of statistical models for repeated measures data.
- It has long played a significant role in empirical research within the developmental sciences, particularly in studying between-individual differences and within-individual patterns of change over time.

# Introduction (cont'd)

- We propose to apply this model in SAE settings to get a model which borrows strength across both small areas and over time by incorporating simultaneously the effects of areas and time interaction.
- This model accounts for repeated surveys, group individuals and random effects variation. The estimation is discussed with a likelihood based approach and a simulation study is conducted.

# The model formulation (cont'd)

- We consider repeated measurements on variable of interest  $y$  for  $p$  time points,  $t_1, \dots, t_p$  from the finite population  $U$  of size  $N$  partitioned into  $m$  disjoint subpopulations or domains  $U_1, \dots, U_m$  called *small areas* of sizes  $N_i, i = 1, \dots, m$  such that  $\sum_{i=1}^m N_i = N$ .
- We also assume that in every area, there are  $k$  different groups of units of size  $N_{ig}$  for group  $g$  such that  $\sum_i^m \sum_{g=1}^k N_{ig} = N$ .
- We draw a sample of size  $n$  in all small areas such that the sample of size  $n_i$  is observed in area  $i$  and  $\sum_i^m \sum_{g=1}^k n_{ig} = n$  and we suppose that we have auxiliary data  $\mathbf{x}_{ij}$  of  $r$  variables (covariates) available for each population unit  $j$  in all  $m$  small areas.

# The model formulation (cont'd)

- The model at Small Area level is given by

$$\begin{aligned} \mathbf{Y}_i &= \mathbf{A}\mathbf{B}_i\mathbf{C}_i + \mathbf{1}\gamma'\mathbf{X}_i + \mathbf{1}\mathbf{u}'_i + \mathbf{E}_i, \\ \mathbf{u}_i &\sim \mathcal{N}_{N_i}(\mathbf{0}, \sigma_u^2\mathbf{I}), \\ \mathbf{E}_i &\sim \mathcal{N}_{p, N_i}(\mathbf{0}, \sigma_e^2\mathbf{I}, \mathbf{I}_{N_i}), \end{aligned} \quad (1)$$

where  $\mathbf{A}$  and  $\mathbf{C}_i$  are respectively *within-individual* and *between-individual* design matrices for fixed effects given by

$$\mathbf{A} = \begin{pmatrix} 1 & t_1 & \cdots & t_1^{q-1} \\ 1 & t_2 & \cdots & t_2^{q-1} \\ \vdots & \vdots & \cdots & \vdots \\ 1 & t_p & \cdots & t_p^{q-1} \end{pmatrix}, \mathbf{C}_i = \begin{pmatrix} 1 & \cdots & 1 & 0 & \cdots & 0 & 0 & \cdots & 0 \\ 0 & \cdots & 0 & 1 & \cdots & 1 & 0 & \cdots & 0 \\ \vdots & \cdots & \vdots & \vdots & \cdots & \vdots & \vdots & \cdots & \vdots \\ 0 & \cdots & 0 & 0 & \cdots & 0 & 1 & \cdots & 1 \end{pmatrix}$$

# The model formulation (cont'd)

- The corresponding model at population level for all small areas can be expressed as

$$\underbrace{\mathbf{Y}}_{p \times N} = \underbrace{\mathbf{A}}_{p \times q} \underbrace{\mathbf{B}}_{q \times mk} \underbrace{\mathbf{C}}_{mk \times N} + \underbrace{\mathbf{1}\gamma'[\mathbf{l}_r : \mathbf{l}_r : \cdots : \mathbf{l}_r]}_{p \times mr} \underbrace{\mathbf{X}}_{mr \times N} + \underbrace{\mathbf{1}}_{p \times 1} \underbrace{\mathbf{u}'}_{1 \times N} + \underbrace{\mathbf{E}}_{p \times N}$$

or

$$\mathbf{Y} = \mathbf{ABC} + \mathbf{1}\gamma'\mathbf{DX} + \mathbf{1u}' + \mathbf{E}, \quad (2)$$

for  $\mathbf{D} = [\mathbf{l}_r : \mathbf{l}_r : \cdots : \mathbf{l}_r]$

# Estimation of model parameters

- In order to transform (2) to a model which is easier to estimate, we transform the design matrix  $\mathbf{A}$  into a new matrix  $\mathbf{A}_1$  with two parts  $\mathbf{A}_1 = [\mathbf{1} : \mathbf{H}]$  and the parameter matrix into a new matrix  $\Xi = [\xi_1 : \Xi_2]$  conformably such that

$$\mathcal{C}(\mathbf{A}) = \mathcal{C}(\mathbf{1}) \oplus \mathcal{C}(\mathbf{H}) \text{ with } \mathcal{C}(\mathbf{H}) = \mathcal{C}(\mathbf{1})^\perp \cap \mathcal{C}(\mathbf{A})$$

- One way of this transformation is given below

$$\mathbf{A} = \begin{pmatrix} 1 & t_1 & \cdots & t_1^{q-1} \\ 1 & t_2 & \cdots & t_2^{q-1} \\ \vdots & \vdots & \cdots & \vdots \\ 1 & t_p & \cdots & t_p^{q-1} \end{pmatrix} \longrightarrow \mathbf{A}_1 = \begin{pmatrix} 1 & t_1 - \bar{t} & \cdots & t_1^{q-1} - \overline{t^{q-1}} \\ 1 & t_2 - \bar{t} & \cdots & t_2^{q-1} - \overline{t^{q-1}} \\ \vdots & \vdots & \cdots & \vdots \\ 1 & t_p - \bar{t} & \cdots & t_p^{q-1} - \overline{t^{q-1}} \end{pmatrix}$$



# Estimation of model parameters (cont'd)

- We come up with the model

$$\mathbf{Y} = \mathbf{1}\xi_1'\mathbf{C} + \mathbf{H}\Xi_2\mathbf{C} + \mathbf{1}\gamma'\mathbf{DX} + \mathbf{1}\mathbf{u}' + \mathbf{E}$$

and make a one-to-one transformation

$$\begin{pmatrix} \mathbf{1}'\mathbf{Y} \\ \mathbf{H}'\mathbf{Y} \\ \mathbf{A}^o'\mathbf{Y} \end{pmatrix} = \begin{pmatrix} p\xi_1'\mathbf{C} + p\gamma'\mathbf{DX} + p\mathbf{u}' + \mathbf{1}'\mathbf{E} \\ \mathbf{H}'\mathbf{H}\Xi_2\mathbf{C} + \mathbf{H}'\mathbf{E} \\ \mathbf{A}^o'\mathbf{E} \end{pmatrix},$$

where  $\mathbf{A}^o$  for a matrix  $\mathbf{A}$  is such that  $\mathbf{A}^o'\mathbf{A} = \mathbf{0}$  and  $\mathcal{C}(\mathbf{A}^o) = \mathcal{C}(\mathbf{A})^\perp$ .

# Estimation of model parameters (cont'd)

After calculation, the maximum likelihood estimators are given by

$$\begin{aligned}\widehat{\Xi}_2 &= (\mathbf{H}'\mathbf{H})^{-1} \mathbf{H}'\mathbf{Y}\mathbf{C}'(\mathbf{C}\mathbf{C}')^{-1} + (\mathbf{H}'\mathbf{H})^{\circ} \mathbf{T}_1 + \mathbf{H}'\mathbf{H}\mathbf{T}_2(\mathbf{C}\mathbf{C}')^{\circ'} \\ \widehat{\gamma}' &= \frac{1}{\rho} \left[ \mathbf{1}'\mathbf{Y}\mathbf{X}'\mathbf{D}' - \mathbf{1}'\mathbf{Y}\mathbf{C}'(\mathbf{C}\mathbf{C}')^{-1}\mathbf{C}\mathbf{X}'\mathbf{D}' - \rho\mathbf{T}_3(\mathbf{C}\mathbf{C}')^{\circ} \mathbf{C}\mathbf{X}'\mathbf{D}' \right] \\ &\quad \times \left[ \mathbf{D}\mathbf{X}\mathbf{X}'\mathbf{D}' - \mathbf{D}\mathbf{X}\mathbf{C}'(\mathbf{C}\mathbf{C}')^{-1}\mathbf{C} \right]^{-1} \\ \widehat{\xi}'_1 &= \left( \frac{1}{\rho} \mathbf{1}'\mathbf{Y} - \widehat{\gamma}'\mathbf{D}\mathbf{X} \right) \mathbf{C}'(\mathbf{C}\mathbf{C}')^{-1} + \mathbf{T}(\mathbf{C}\mathbf{C}')^{\circ}\end{aligned}$$

for some matrices  $\mathbf{T}$ ,  $\mathbf{T}_1$ ,  $\mathbf{T}_2$  and  $\mathbf{T}_3$  of proper sizes.

# Estimation of model parameters (cont'd)

- Once  $\widehat{\xi}'_1$  and  $\widehat{\Xi}_2$  are obtained, we can then find the parameter matrix  $\mathbf{B}$  by solving the linear system

$$\mathbf{1}\widehat{\xi}'_1\mathbf{C} + \mathbf{H}\widehat{\Xi}_2\mathbf{C} = \mathbf{A}\widehat{\mathbf{B}}\mathbf{C}.$$

Since, the matrices  $\mathbf{A}$  and  $\mathbf{C}$  are of full rank, then

$$\widehat{\mathbf{B}} = (\mathbf{A}'\mathbf{A})^{-1}\mathbf{A}'\left(\mathbf{1}\widehat{\xi}'_1\mathbf{C} + \mathbf{H}\widehat{\Xi}_2\mathbf{C}\right)\mathbf{C}'(\mathbf{C}\mathbf{C}')^{-1}.$$

# Estimation of model parameters (cont'd)

- Given the covariance structure of  $\mathbf{Y}$

$$\boldsymbol{\Sigma} = \mathbf{1}\boldsymbol{\Sigma}_u\mathbf{1}' + \boldsymbol{\Sigma}_e = m\sigma_u^2\mathbf{1}\mathbf{1}' + \sigma_e^2\mathbf{I}_p,$$

and its inverse

$$\boldsymbol{\Sigma}^{-1} = \frac{1}{\sigma_e^2} \left( \mathbf{I}_p - \frac{m\sigma_u^2}{mp\sigma_u^2 + \sigma_e^2} \mathbf{1}\mathbf{1}' \right).$$

- We find the maximum likelihood estimator of the variance component expressed by

$$\hat{\sigma}_u^2 = \frac{\text{tr}\{\mathbf{1}\mathbf{1}'\mathbf{W}\} - Np\sigma_e^2}{Nmp^2},$$

where

$$\mathbf{W} = (\mathbf{Y} - \mathbf{ABC} - \mathbf{1}\boldsymbol{\gamma}'\mathbf{DX})(\mathbf{Y} - \mathbf{ABC} - \mathbf{1}\boldsymbol{\gamma}'\mathbf{DX})'.$$

# Prediction of random effects

- Under the theory of linear model and normal distribution, the best linear predictor of  $u$  that minimizes the mean square error is the conditional mean  $E[\mathbf{u}|\mathbf{Y}]$  given by

$$E[\mathbf{u}|\mathbf{Y}] = E[\mathbf{u}] + \text{Cov}(\mathbf{u}', \mathbf{Y})\text{Cov}^{-1}(\mathbf{Y})(\mathbf{Y} - E[\mathbf{Y}]).$$

- Thus,

$$\begin{aligned}\hat{\mathbf{u}} &= \hat{\sigma}_u^2 \mathbf{1}' \hat{\Sigma}^{-1} (\mathbf{Y} - \mathbf{A}\hat{\mathbf{B}}\mathbf{C} - \mathbf{1}\hat{\gamma}'\mathbf{D}'\mathbf{X}) \\ &= \frac{\hat{\sigma}_u^2}{mp\hat{\sigma}_u^2 + \sigma_e^2} \mathbf{1}' (\mathbf{Y} - \mathbf{A}\hat{\mathbf{B}}\mathbf{C} - \mathbf{1}\hat{\gamma}'\mathbf{D}'\mathbf{X})\end{aligned}$$

# Simulation study Example

We consider 6 small areas and draw a sample with the following sample sizes.

Table : Sample sizes

Area	Group 1	Group 2	Total
1	$n_{11}=52$	$n_{12}=48$	$n_1=100$
2	$n_{21}=60$	$n_{22}=60$	$n_2=120$
3	$n_{31}=30$	$n_{32}=40$	$n_3=70$
4	$n_{41}=46$	$n_{42}=22$	$n_4=68$
5	$n_{51}=65$	$n_{52}=65$	$n_5=130$
6	$n_{61}=50$	$n_{62}=62$	$n_6=112$
$m=6$	$g_1=303$	$g_2=297$	$n=600$

We assume  $p = 4$  and  $r = 3$ .

# Simulation study Example (cont'd)

The design matrices are

$$\mathbf{A} = \begin{pmatrix} 1 & 1 \\ 1 & 2 \\ 1 & 3 \\ 1 & 4 \end{pmatrix} \rightarrow \mathbf{H} = \begin{pmatrix} -1.5 \\ -0.5 \\ 0.5 \\ 1.5 \end{pmatrix},$$

$$\mathbf{C} = \begin{pmatrix} \mathbf{C}_1 & \cdots & \mathbf{0} \\ \mathbf{0} & & \mathbf{C}_6 \end{pmatrix} \quad \text{for} \quad \mathbf{C}_i = \left( \mathbf{1}'_{n_{i1}} \otimes \begin{pmatrix} 1 \\ 0 \end{pmatrix} : \mathbf{1}'_{n_{i2}} \otimes \begin{pmatrix} 0 \\ 1 \end{pmatrix} \right),$$

$i = 1, \dots, 6;$

# Simulation study Example (cont'd)

The parameter matrices are

$$\xi'_1 = (20 \ 21 \ 22 \ 23 \ 24 \ 25 \ 26 \ 27 \ 28 \ 29 \ 30 \ 31),$$

$$\Xi_2 = (1 \ 2 \ 3 \ 4 \ 5 \ 6 \ 7 \ 8 \ 9 \ 10 \ 11 \ 12),$$

$$\begin{aligned} \mathbf{B} &= \mathbf{A}^{-1} \left( \mathbf{1} \xi'_1 \mathbf{C} + \mathbf{H} \Xi_2 \mathbf{C} \right) \mathbf{C}^{-1} \\ &= \begin{pmatrix} 17.5 & 16 & 14.5 & 13 & 11.5 & 10 & 8.5 & 7 & 5.5 & 4 & 2.5 & 1 \\ 1 & 2 & 3 & 4 & 5 & 6 & 7 & 8 & 9 & 10 & 11 & 12 \end{pmatrix}, \end{aligned}$$

and

$$\gamma = \begin{pmatrix} 1 \\ 2 \\ 3 \end{pmatrix}, \sigma_u^2 = 5, \sigma_e^2 = 6.$$



# Simulation study Example (cont'd)

Then, the data are generated from

$$\mathbf{Y} \sim N_{p,n}(\mathbf{ABC} + \mathbf{1}\gamma'\mathbf{DX}, \boldsymbol{\Sigma}, \mathbf{I}_n),$$

where the matrix of covariates  $\mathbf{X}$  is generated with random elements.  
The following MLEs are obtained:

$$\widehat{\boldsymbol{\xi}}_1' = \begin{pmatrix} 20.2534 & 21.6548 & 22.5961 & 23.6486 & 24.4233 & 25.0374 & 25.998 \\ 28.5361 & 29.9077 & 30.3292 & 31.1121 \end{pmatrix}$$

$$\widehat{\boldsymbol{\Xi}}_2 = \begin{pmatrix} 1.1151 & 2.0824 & 3.0320 & 3.6376 & 4.6384 & 5.7882 & 7.0238 & 7.877 \\ 9.0386 & 10.1256 & 10.8561 & 11.9422 \end{pmatrix}$$

# Simulation study Example (cont'd)

$$\hat{\mathbf{B}} = \begin{pmatrix} 17.4657 & 17.3902 & 14.0748 & 14.5546 & 12.8274 & 10.5669 & 8.4390 \\ 1.1151 & 2.0824 & 3.0320 & 3.6376 & 4.6384 & 5.7882 & 7.0238 \\ 5.9397 & 4.5936 & 3.1890 & 1.2566 & & & \\ 9.0386 & 10.1256 & 10.8561 & 11.9422 & & & \end{pmatrix}$$







$$\hat{\sigma}_u^2 = 5.0061, \quad \hat{\gamma} = \begin{pmatrix} 1.0093 \\ 1.9501 \\ 3.0469 \end{pmatrix}, \quad \mathbf{ABC} = \begin{pmatrix} 18.5 & \dots & 18 & \dots & 13 \\ 19.5 & \dots & 20 & \dots & 25 \\ 20.5 & \dots & 22 & \dots & 37 \\ 21.5 & \dots & 24 & \dots & 49 \end{pmatrix}$$

$$\widehat{\mathbf{ABC}} = \begin{pmatrix} 18.5808 & \dots & 18.4726 & \dots & 13.1988 \\ 19.6959 & \dots & 20.5550 & \dots & 25.1410 \\ 20.8110 & \dots & 22.6373 & \dots & 37.0833 \\ 21.9261 & \dots & 24.7197 & \dots & 49.0255 \end{pmatrix}$$







# Further research

- After obtaining all unknown parameters, then we can find directly the target small area characteristics of interest such as the small area totals and small area means
- In further research, we want to test the efficiency, the distribution and all properties of the estimators
- We wish also to study the possible time correlation

# Some references

-  Bai Peng, *Exact distribution of MLE of covariance matrix in GMANOVA-MANOVA model*, J. Science in China, 2005
-  Battese, G.E, R.M. and W.A. Fuller, *An error-components model for prediction of county crop areas using survey and satellite data*, American Statistical Association, 1988.
-  Danny Pfeffermann *Small Area Estimation-New Developments and Directions*. J. International Statistical Review, 2002.
-  G. Datta, P. Lahiri, T. Maiti, K. Lu, *Hierarchical Bayes estimation of unemployment rates for the states of the US*, Journal of the American Statistical Association 94 (1999)
-  G.K. Robinson, *That BLUP Is a Good Thing: The estimation of Random Effects*, Statistical Science, Vol. 6, 1991
-  J.N.K. Rao, *Small Area Estimation*. Willey, 2003.

# Some references

-  T. Kollo and D. von Rosen, *Advanced Multivariate Statistics with matrices*. Springer, 2005.
-  Kari Nissinen, *Small Area Estimation with Linear Mixed Models from Unit-level panel and Rotating panel data*. PhD Thesis, Jyväskylä University, 2009..
-  M. Ghosh and J.N.K. Rao, *Small Area Estimation: An Appraisal*, J. Statistical Science, 1994.
-  R. Chambers and R. G. Clark, *An introduction to Model-Based Survey Sampling with Applications*. Oxford, 2012.
-  Robb J. Muirhead, *Aspects of Multivariate Statistical theory* , Wiley 2005.
-  Tatsuya Kubokawa and Muni S. Srivastava, *Prediction in Multivariate Mixed Linear Models*, J. Japan Statist. Soc, 2003

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